

A list of useful things

- Addition rule: $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
- Conditional probability: $P(A | B) = \frac{P(A \cap B)}{P(B)}$.
- Events A and B are independent if $P(A \cap B) = P(A) \cdot P(B)$, or equivalently $P(A | B) = P(A)$, or $P(B | A) = P(B)$

Probability distributions:

	Discrete	Continuous
Probability density function (pdf)	$p(x) \geq 0, \sum_x p(x) = 1$	$f(x) \geq 0, \int_R f(x) dx = 1$
Cumulative distribution function (cdf)	$F(x) = P(X \leq x) = \sum_{y \leq x} p(y)$	$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$
Expected value	$\mu = E(X) = \sum_x xp(x)$	$\mu = E(X) = \int_R xf(x) dx$
Variance	$\sigma^2 = E(X^2) - \mu^2 = \sum_x x^2 p(x) - \mu^2$	$\sigma^2 = E(X^2) - \mu^2 = \int_R x^2 f(x) dx - \mu^2$
Standard deviation	$\sigma = \sqrt{\sigma^2}$	$\sigma = \sqrt{\sigma^2}$
Moment generating function (MGF)	$M(t) = E(e^{tX}) = \sum_x e^{tx} p(x)$	$M(t) = E(e^{tX}) = \int_R e^{tx} f(x) dx$

Computing with the MGF

$E(X) = M'(0)$

$E(X^2) = M''(0), \dots$, in general $E(X^n) = M^{(n)}(0)$

whenever the derivatives of the moment generating function exist.

$\sigma^2 = M''(0) - [M'(0)]^2$.

Binomial(n,p)	$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$	$E(X) = np$	$V(X) = np(1-p)$	$M(t) =$
Hypergeometric	$P(X = k) = \frac{\binom{M}{k} \binom{N-M}{n-k}}{\binom{N}{n}}$	$E(X) = n \frac{M}{N}$	$V(X) = n \frac{M}{N} \frac{N-M}{N} \frac{N-n}{N-1}$	$M(t) =$
Poisson(λ)	$P(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$	$E(X) = \lambda$	$V(X) = \lambda$	$M(t) =$
Uniform(a,b)	$f(x) = \frac{1}{b-a}, a \leq x \leq b$	$E(X) = \frac{a+b}{2}$	$V(X) = \frac{(b-a)^2}{12}$	$M(t) =$

Exponential(λ)	$f(x) = \frac{1}{\lambda} e^{-\frac{x}{\lambda}}, x > 0$	$E(X) = \lambda$	$V(X) = \lambda^2$	$M(t) =$
Normal(μ, σ)	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$	$E(X) = \mu$	$V(X) = \sigma^2$	$M(t) =$

Note: In the above, “exp” stands for exponential

Joint Probability Distributions

	Discrete	Continuous
Joint pdf	$p(x, y) \geq 0, \sum_x \sum_y p(x, y) = 1$	$f(x, y) \geq 0, \iint_D f(x, y) dx dy = 1$
Marginal of X	$P(X = x) = f_X(x) = \sum_y p(x, y)$	$f_X(x) = \int f(x, y) dy$
Marginal of Y	$P(Y = y) = f_Y(y) = \sum_x p(x, y)$	$f_Y(y) = \int f(x, y) dx$
Independence of X and Y	$p(x, y) = f_X(x) f_Y(y)$	$p(x, y) = f_X(x) f_Y(y)$

Linear Combination of Random Variables

Let X and Y be random variables and a, b real numbers.

$$E(aX + bY) = aE(X) + bE(Y)$$

If X and Y are **independent** then $V(aX + bY) = a^2V(X) + b^2V(Y)$.

If X and Y are independent and u, v are functions, then $E(u(X)v(Y)) = E(u(X))E(v(Y))$.

- If X is binomial(n, p) and Y is binomial (m, p), and X and Y are independent, then $X+Y$ is Binomial ($n + m, p$).
- If X and Y are independent Poisson random variables with means λ_1, λ_2 respectively, then $X+Y$ is Poisson($\lambda_1 + \lambda_2$).
- If X and Y are independent normally distributed random variables, $N(\mu_1, \sigma_1^2)$, and $N(\mu_2, \sigma_2^2)$ respectively, then $aX + bY$ is $N(a\mu_1 + b\mu_2, a^2\sigma_1^2 + b^2\sigma_2^2)$.

- More general, if X_1, X_2, \dots, X_n are independent normal random variables,

$$X_i \sim N(\mu_i, \sigma_i^2), \text{ then } \sum_{i=1}^n a_i X_i \sim N(\mu, \sigma^2), \mu = \sum_{i=1}^n a_i \mu_i, \sigma^2 = \sum_{i=1}^n a_i^2 \sigma_i^2$$

- If X is $N(\mu, \sigma^2)$ then $Z = \frac{X - \mu}{\sigma}$ is standard normal.
- If Z is standard normal, then Z^2 is $\chi^2(1)$ Chi square with one degree of freedom.
- If Z_1, \dots, Z_n is a random sample from the standard normal distribution, then

$$W = Z_1^2 + Z_2^2 + \dots + Z_n^2 \text{ is } \chi^2(n)$$

○ If X_1, \dots, X_n are independent, normally distributed, $N(\mu, \sigma^2)$ and $S^2 = \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{n-1}$, then $\frac{(n-1)S^2}{\sigma^2}$ is $\chi^2(n-1)$.

○ If X_1, \dots, X_n are independent, normally distributed, $N(\mu, \sigma^2)$ then $T = \frac{\bar{X} - \mu}{S/\sqrt{n}}$ has Student t-distribution with $n-1$ degrees of freedom.

○ If X_1, \dots, X_n are independent, normally distributed, $N(\mu_i, \sigma_i^2)$ then

$$W = \sum_{i=1}^n \frac{(X_i - \mu_i)^2}{\sigma_i^2} \text{ is } \chi^2(n).$$

The Central Limit Theorem

Let X_1, \dots, X_n be a random sample from a distribution with mean μ and standard deviation

σ . Then $\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$ is approximately $N(\mu, \frac{\sigma^2}{n})$.

In particular the result applies for a sum of a random sample: $T = \sum_{i=1}^n X_i$ is approximately normally distributed with mean $E(T) = n\mu$ and variance $Var(T) = n\sigma^2$.